# Estimating an Equilibrium Model 

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## Motivation

- In low-income countries, schooling does not equal learning.
- In rural Pakistan, $25 \%$ of students do not experience test score gains from year to year.
- Policies focus on improving "vertical" school quality (teacher training, improved inputs) and have small or moderate effects.
- Mismatch between instruction level and student ability is important.
- Particularly true in low-income countries.
- Know little about whether misallocation of instructional level affects learning and how it responds to market structure.
- Private schooling share is high in low-income countries.


## This Paper I

- To motivate the importance of instructional match, provide evidence that private school entry increases within-school inequality in test scores in private schools.
- Develop and estimate a novel model in which schools choose instruction level and vertical quality.
- Profit-maximizing private schools take into account different students' differential responsiveness to quality leading to a Spence Distortion.
- Apply this general framework to education markets in Pakistan.


## This Paper II

- Estimate the distribution of private schools' instruction levels.
- The average private school strongly caters to more advantaged students.
- Provides a structural estimate of the importance of match-specific quality.
- Moving a student from their worst to best match school increases test scores by $\approx 1$ year of test score gains.
- Model predicts that private school entry leads to greater inequality between more and less disadvantaged students' learning within a school.
- Confirm this is the case in difference-in-differences regressions.
- Calculate partial equilibrium counterfactuals where instructional levels are set to maximize learning.
- Greatly reduces inequality in learning and moderately increases average learning.
- Low responsiveness to quality in enrollment decisions limits gains in counterfactuals.


## Context: Pakistan

- In rural Pakistan, 35\% of students are enrolled in unregulated private schools.
- The typical school has 1 or less teachers per grade.
- The mean private school in the data costs $5 \%$ of per capita annual income.
- Villages are closed educational markets. $\Rightarrow$ Why is this important?
- Average village has $\approx 3$ private schools.


## Data

- LEAPS Pakistan data consists of a random sample of 112 villages in the province of Punjab.
- Three key surveys (each conducted every year from 2004-2007):
- Geo-coded survey of universe of schools.
- Survey of children in the schools (child survey), including low-stakes test scores.
- Geo-coded survey of a sub-sample of households in the village (household survey).
- Importantly, some overlap in the household and child survey.


## Market Structure and Test Scores

To provide evidence that school quality is different for different students,

- Estimate the effects of private school entry on the average student's test scores in public and private schools.
- Then, estimate the effect of private school entry on inequality in test scores in the public and private sector.
- Drivers of Exit and Entry


## Effect of Competition on Average Test Scores

|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Math |  | English |  | Urdu |  | Mean |  |
|  | Private | Public | Private | Public | Private | Public | Private | Public |
| num_privt | -0.078** | 0.006 | -0.039 | -0.010 | -0.010 | -0.031 | -0.038 | -0.012 |
|  | (0.031) | (0.045) | (0.046) | (0.037) | (0.036) | (0.030) | (0.034) | (0.036) |
| Num Gov. School Control | Y | Y | Y | Y | Y | Y | Y | Y |
| Report Card-Year FE | Y | Y | Y | Y | Y | Y | Y | Y |
| Lagged Test Scores | Y | Y | Y | Y | Y | Y | Y | Y |
| Grade FE | Y | Y | Y | Y | Y | Y | Y | Y |
| Village FE | Y | Y | Y | Y | Y | Y | Y | Y |
| Number of observations | 6,788 | 13,938 | 6,788 | 13,938 | 6,788 | 13,938 | 6,788 | 13,938 |
| Clusters | 108 | 112 | 108 | 112 | 108 | 112 | 108 | 112 |
| Adjusted $\mathrm{R}^{2}$ | 0.534 | 0.565 | 0.550 | 0.571 | 0.611 | 0.606 | 0.655 | 0.671 |

No effect on the test scores of the average student in either sector.

## Effect of Competition on Inequality

|  | Variance Math |  | (3) | (4) | (5) | (6) | (7) | (8) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Variance English |  | Variance Urdu |  | Variance Mean |  |
|  | Private | Public | Private | Public | Private | Public | Private | Public |
| num_privt | 0.085*** | 0.022 | 0.096*** | 0.047 | 0.098*** | 0.043 | 0.077*** | 0.034 |
|  | (0.026) | (0.025) | (0.032) | (0.031) | (0.033) | (0.030) | (0.023) | (0.025) |
| Num. Gov. School Control | Y | Y | Y | Y | Y | Y | Y | Y |
| Report Card-Year FE | Y | Y | Y | Y | Y | Y | Y | Y |
| Village FE | Y | Y | Y | Y | Y | Y | Y | Y |
| Number of observations | 426 | 447 | 426 | 447 | 426 | 447 | 426 | 447 |
| Clusters | 108 | 112 | 108 | 112 | 108 | 112 | 108 | 112 |
| Adjusted $\mathrm{R}^{2}$ | 0.292 | 0.400 | 0.248 | 0.362 | 0.242 | 0.311 | 0.322 | 0.412 |

Note: variance is calculated at the village-year level.

- Null effect on the average student masks significant heterogeneous effects in the private sector.


## Effect of Competition on Inequality

Figure: Effect of Entry on Village-Level Variance in Private-Sector Test Scores

inequality $_{v t}=\rho_{0}+\gamma_{1}$ event $_{v, s-3}+\gamma_{2}$ event $_{v, s-2}+\gamma_{3}$ event $_{v, 0}$

$$
+\gamma_{4} \text { event }_{v, s+1}+\gamma_{5} \text { event }_{v, s+2}+\alpha_{t}+\psi_{v}+\Gamma \mathbf{X}_{\mathrm{vt}}+\epsilon_{v t}
$$

## What Drives the Increase in Private Sector Inequality?

|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Variance Math |  | Variance English |  | Variance Urdu |  | Variance Mean |  |
|  | No <br> Switchers | School-Level Variance | No Switchers | School-Level Variance | No Switchers | School-Level Variance | No Switchers | School-Level Variance |
| num_privt | 0.099** | 0.089*** | 0.078* | 0.063** | 0.130** | 0.049 | 0.089*** | 0.060** |
|  | (0.038) | (0.032) | (0.041) | (0.027) | (0.050) | (0.036) | (0.033) | (0.024) |
| Num Gov. School Control | Y | Y | Y | Y | Y | Y | Y | Y |
| Report Card-Year FE | Y | Y | Y | Y | Y | Y | Y | Y |
| School FE | N | Y | N | Y | N | Y | N | Y |
| Village FE | Y | N | Y | N | Y | N | Y | N |
| Peer Controls | Y | N | Y | N | Y | N | Y | N |
| Student-Teacher Ratio Control | Y | N | Y | N | Y | N | Y | N |
| Number of observations | 405 | 802 | 405 | 802 | 405 | 802 | 405 | 802 |
| Clusters | 106 | 103 | 106 | 103 | 106 | 103 | 106 | 103 |
| Adjusted R ${ }^{2}$ | 0.235 | 0.181 | 0.211 | 0.134 | 0.222 | 0.268 | 0.279 | 0.224 |

- Cannot be explained by changes in the schools that students attend or the composition of the private sector.
- Variance in test scores increases within schools.
- Robust to standard controls for peer effects.


## Evidence on Mechanisms

|  | $(1)$ <br> Any Bonuses | $(2)$ <br> Bonuses for <br> Exam Performance | $(3)$ <br> Teacher Attendance | (4) <br> Share Getting <br> Bonuses for Exams |
| :--- | :---: | :---: | :---: | :---: |
| Num. Pri. Schools | 0.026 | $0.082^{* * *}$ | $-0.052^{* *}$ | $0.037^{* * *}$ |
|  | $(0.025)$ | $(0.027)$ | $(0.022)$ | $(0.013)$ |
| Mean Dep. Var. | 0.373 | 0.165 | 0.165 | 0.035 |
| Number of observations | 1,151 | 1,151 | 1,151 | 1,145 |
| Clusters | 108 | 108 | 108 | 108 |
| Adjusted $\mathrm{R}^{2}$ | 0.304 | 0.143 | 0.210 | 0.073 |

- Private schools become more likely to give teacher bonuses based on students' test scores instead of teacher attendance.


## Model: Environment

- Students choose schools or the outside option to maximize their utility based on school characteristics, including quality.
- Students differ in their responsiveness to schools' characteristics.
- Students' learning in a school depends on both vertical quality and instructional match.
- Vertical quality: better facilities, better trained teacher, lower teacher absence.
- Instructional match: a student who cannot do arithmetic does not benefit from learning calculus.
- Formally,
- Poorer Students: optimal instructional level is 0 on the Hotelling line.
- Richer Students: optimal instructional level is 1 on the Hotelling line.
- $N$ schools choose their characteristics to maximize their profits.


## Student's Problem

A student $i$ of type $z$ chooses a school $j$ or not to enroll to maximize

$$
u_{i j z}=\delta_{z} V A_{j z}+\Gamma_{z} X_{i j}+\xi_{j}+\epsilon_{i j}
$$

where

$$
V A_{j, p o o r}=v_{j}-\beta\left(0-h_{j}\right)^{2}
$$

and

$$
V A_{j, r i c h}=v_{j}-\beta\left(1-h_{j}\right)^{2}
$$

- $h_{j}$ : school $j$ 's choice of instructional level on the unit line.
- $X_{i j}$ : Characteristics of school $j$ and student $i$.
- $\xi_{j}$ : School quality unobserved to the econometrician.
- $\epsilon_{i j}$ : type 1 extreme value error.


## School's Problem

A school chooses its instructional level $h_{j} \in[0,1]$, its vertical quality $v_{j}$, and its price, fee ${ }_{j}$, to maximize

$$
\pi_{j}=f e e_{j} \times s_{j}\left(v_{j}, h_{j}, f e e_{j}, X_{j}, \xi_{j}\right)-c\left(v_{j}, h_{j}, f e e_{j}, X_{j}, \xi_{j}\right)
$$

- $s_{j}$ : the number of students who attend school $j$.
- $c$ : the cost of providing education, which depends on both the school's quality and the number of students.


## How Can Competition Lead to Increased Inequality?

A simple example illustrates how competition can lead to increased inequality:

- Assume instructional match is the only characteristic that matters $\left(\Gamma_{z}=0, \xi_{j}=0, v_{j}=0\right)$ and that price is fixed.
- There is one free, non-private option that gives utility $u_{o}$.
- Equal number of rich and poor students.
- $\delta_{\text {rich }}$ is $\infty$ : Richer students always maximize learning.
- $\delta_{\text {poor }}<\delta_{\text {rich }}$ : Poorer students respond positively to expected learning.
- There are $N \in\{1,2\}$ private schools.


## Equilibrium With 1 School

## Prediction

For $N=1$, there is a unique equilibrium where the school places at $h^{*}=1-\left(-u_{o}\right)^{1 / 2}$.


## Equilibrium With 2 Schools

Prediction
For $N=2$, if $\delta_{\text {poor }}$ is sufficiently small, the unique equilibrium is $(1,1)$.


## Key Intuition

- Spence distortion: oligopolistic schools respond to the marginal consumer rather than inframarginal consumers when they choose their quality.
- Advantaged students are more responsive to quality when they make their enrollment decisions.
- Increasing competition makes these students "more marginal."
- Schools' choices of quality will be very different from the social planner's since the social planner cares about inframarginal students.
- In other words, schools may respond to competitive incentives by competing to have the most advanced instructional level, reducing learning.


## Mapping Model to Data

- To map the model to data, we need to measure both student types and school quality.
- First
- Assign student types in two different data sets.
- Child survey: contains test score data.
- Household survey: contains distance to school and school choice data.
- Estimate school quality.
- Then
- Estimate the model.


## Measuring School Quality

Use value-added methodology to estimate a type z's test scores $y_{i t}$ in a school $j$ :

$$
y_{i t}=\tau_{g z, 1} y_{i, t-1}+\tau_{g z, 2} y_{i, t-1}^{2}+\omega_{g z}+\alpha_{z t}+V A_{z j n}+\epsilon_{i t}
$$

where $i$ is a student, $g$ is a grade, $z$ is a type, $t$ is a year, and $n$ is the number of competitors in the market place.

- $V A_{z j n}$ : school by type by number of competitors fixed effect.
- $\tau_{g z, 1}, \tau_{g z, 2}$ : grade by type specific coefficients.
- $\omega_{g z}$ : grade by type fixed effect.
- $\alpha_{z t}$ : type by year fixed effect.


## Measuring Types in the (In-School) Child Survey

- Factor analysis of normalized HH assets and their interactions.
- Predict the first factor.
- Separate students into those above and below a data-driven cut-off for the private school population.


## Data-Driven Cut-off

Estimate

$$
y_{i t}=\tau_{g z, 1} y_{i, t-1}+\tau_{g z, 2} y_{i, t-1}^{2}+\omega_{g z}+\alpha_{z t}+\phi_{j n}+\sum_{j n} \kappa_{j n} \mathbb{1}_{i}^{r i c h}+\epsilon_{i t}
$$

- Choose cut-off that maximizes F-statistic for $\kappa_{j n}$.
- Selects 55th percentile.


## Types in the Household Data

- Child survey does not include distance measures.
- Household survey does not include test scores.
- To calculate types in household survey, take advantage of overlap in data.
- Lasso logistic regression of type in household survey on normalized HH assets for students in the overlapping sample.
- Predict probability of being an advantaged type $p_{H}$ in full household survey.


## Estimating the Model

Recall

$$
u_{i j z t}=\delta_{z} V A_{j z}+\Gamma_{z} X_{i j t}+\xi_{j n}+\epsilon_{i j t}
$$

where

$$
V A_{j, \text { dis }}=v_{j}-\beta\left(0-h_{j}\right)^{2}
$$

and

$$
V A_{j, a d}=v_{j}-\beta\left(1-h_{j}\right)^{2} .
$$

To estimate these parameters:

1. Estimate the determinants of school choice (demand side).

- Find that $\delta_{\text {poor }}<\delta_{\text {rich }}$.

2. Estimate equilibrium choice of school quality (supply side).

- Find that $h_{j t}$ is increasing in the number of schools.
- Estimate $\beta$.


## Demand Estimation

To estimate the parameters of the utility function

$$
u_{i j z t}=\delta_{z} V A_{j z n}+\Gamma_{z} X_{i j t}+\xi_{j n}+\epsilon_{i j t},
$$

first, define

$$
\zeta_{j n}=\xi_{j n}+\Gamma_{z}^{\text {school }} X_{j n}^{\text {school }}
$$

Then, estimate $\left\{\Gamma_{z}^{\text {indiv }}, \delta_{\text {poor }}, \delta_{\text {rich }}, \zeta_{j n}\right\}$ by maximizing

$$
\sum_{i j t} \mathbb{1}_{i j t} \log \left(p_{i j t}\right)
$$

where

$$
\begin{aligned}
& p_{i j t}=P\left(\text { type }_{i}=r i c h\right) \frac{e^{\delta_{\text {rich }} V A_{j, \text { rich }, n}+\Gamma_{\text {rich }}^{\text {indiv }}} X_{i j t}^{\text {indiv }}+\zeta_{j n}}{\sum_{k} e^{\delta_{r i c h} V A_{k, \text { rich }, n}+\Gamma_{\text {rich }}^{\text {indi }}} X_{i k t}^{\text {indiv }}+\zeta_{k n}} \\
& +\left(1-P\left(\text { type }_{i}=\text { rich }\right)\right) \frac{e^{\delta_{\text {poor }} V A_{j, p o o r, n}+\Gamma_{\text {poor }}^{\text {indiv }} X_{i j t}^{\text {poor }}+\zeta_{j n}}}{\sum_{k} e^{\delta_{\text {poor }} V A_{k, \text { poor }, n}+\Gamma_{\text {poor }}^{\text {idi }} X_{i k t}^{\text {indiv }}+\zeta_{k n}} .}
\end{aligned}
$$

## Demand Estimation Results

\(\left.$$
\begin{array}{lcc}\hline & \begin{array}{c}(1) \\
\text { Coefficient }\end{array}
$$ \& (2) <br>

Se\end{array}\right]\)| distance $_{i j} \times \mathbb{1}_{\text {poor }}$ | $-1.447^{* * *}$ | 0.076 |
| :--- | :---: | :---: |
| distance $_{i j} \times \mathbb{1}_{\text {rich }}$ | $-0.393^{* * *}$ | 0.050 |
| $V A_{j, \text { poor }, n} \times \mathbb{1}_{\text {poor }}$ | $0.520^{*}$ | 0.283 |
| $V A_{j, \text { rich }, n} \times \mathbb{1}_{\text {rich }}$ | $1.218^{* * *}$ | 0.215 |
| fee $_{j n} \times \mathbb{1}_{\text {rich }}$ | $0.212^{*}$ | 0.114 |
| fee $_{j n}$ | $-4.521^{* * *}$ | 0.197 |

- Do the demand estimates make intuitive sense?


## Supply Estimation: Equilibrium Choice of Horizontal Quality

Recall

$$
\pi_{j}=f e e_{j} \times s_{j}\left(v_{j}, h_{j}, f e e_{j}, X_{j}, \xi_{j}\right)-c\left(v_{j}, h_{j}, f e e_{j}, X_{j}, \xi_{j}\right)
$$

SO

$$
\frac{\partial \pi_{j}}{\partial h_{j}^{*}}=\left(f e e_{j}-\frac{\partial c}{\partial s_{j}}\right) \frac{\partial s_{j}}{\partial h_{j}^{*}}-\frac{\partial c}{\partial h_{j}^{*}}=0
$$

## Equilibrium Choice of Horizontal Quality

If we assume $\frac{\partial c}{\partial h_{j}^{*}}=0$ and $c$ is weakly-concave in share of students
$s$, profit-maximization implies that

$$
\frac{\partial s_{j t}}{\partial h_{j t}}=0
$$

- Assume we observe schools choosing their equilibrium characteristics (fees, vertical quality) in the data.
- Then, we can solve for what equilibrium horizontal quality a private school should choose given its own characteristics and the characteristics of the other schools in the market.


## Estimates of Equilibrium Horizontal Quality



Average horizontal quality is 0.7 , even though $55 \%$ of private school students have an optimal instructional level of 0 .

## Interpreting Estimates

- Spence distortions are important and lead private schools to compete more intensively for wealthy students.
- Driven by the fact $\delta_{\text {rich }}>\delta_{\text {poor }}$.
- From the point of view of a social planner who wants to maximize learning, instructional level is misallocated.
- $h_{j t}$ is positively correlated with number of private schools, consistent with motivating results.


## Identifying $\beta$

Using the fact that

$$
V A_{j, p o o r}=v_{j}-\beta h_{j}^{2}
$$

and

$$
V A_{j, r i c h}=v_{j}-\beta\left(1-h_{j}\right)^{2}
$$

identify $\beta$ from

Key intuition: the relationship between schools' optimal choices of $h_{j t}$ and the difference between school quality for disadvantaged and advantaged students identifies $\beta$.

## Importance of Horizontal Quality

Recall:

$$
V A_{j, p o o r}=v_{j}-\beta h_{j}^{2}
$$

and

$$
V A_{j, r i c h}=v_{j}-\beta\left(1-h_{j}\right)^{2}
$$

- $\beta$ captures the relative importance of instructional match for learning.
- Estimate $\beta=0.36$.
- Interpretation: moving a school's instructional level from 1 to 0 increases a disadvantaged student's mean test score by 0.36 sd.


## Horizontal Quality and Number of Schools



Delivers a prediction that we can test with our reduced-form exit-entry identification strategy.

## Testing the Model Predictions

Test whether increased competition increases inequality between rich and poor students in the same private school:

$$
\begin{aligned}
y_{i t}= & \rho_{0}+\rho_{1} n u m \_p r i_{v t}+\rho_{2} n u m \_p r i_{v t} \times \mathbb{1}_{r i c h}+\eta_{z j}+\alpha_{z t} \\
& +\omega_{g z}+\lambda_{g z} y_{i, t-1}+\phi_{g z} y_{i, t-1}^{2}+\epsilon_{i t}
\end{aligned}
$$

- num_privt: number of private schools in village $v$ and year $t$.
- $\mathbb{1}_{\text {ad }}$ : indicator variable for advantaged.
- $\eta_{z j}$ : school by type fixed effect.
- $\alpha_{z t}$ : type by year fixed effect.
- $\omega_{g z}$ : type by grade fixed effect.
- $\lambda_{g z}, \phi_{g z}$ : grade by type specific coefficient.

Potential confounders:

- Time trends.
- Compositional changes.


## Exit and Entry of Private Schools

|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Math |  | English |  | Urdu |  | Mean |  |
| $\mathbb{1}_{\text {rich }} \times$ num_privt | 0.168*** | 0.145** | 0.103** | 0.093** | 0.090* | 0.066 | 0.121*** | 0.100** |
|  | (0.055) | (0.056) | (0.043) | (0.045) | (0.046) | (0.048) | (0.038) | (0.040) |
| $n u m \_p r i_{v t}$ | -0.132*** | -0.137*** | -0.066 | -0.068 | -0.033 | -0.036 | -0.073* | -0.077** |
|  | (0.036) | (0.038) | (0.056) | (0.056) | (0.043) | (0.041) | (0.037) | (0.035) |
| Num. Gov School Control | Y | Y | Y | Y | Y | Y | Y | Y |
| Peer Controls | N | Y | N | Y | N | Y | N | Y |
| Student-Teacher Ratio Controls | N | Y | N | Y | N | Y | N | Y |
| Lagged Test Score Controls | Y | Y | Y | Y | Y | Y | Y | Y |
| School by Type FE | Y | Y | Y | Y | Y | Y | Y | Y |
| Grade by Type FE | Y | Y | Y | Y | Y | Y | Y | Y |
| Year-Type-Report Card FE | Y | Y | Y | Y | Y | Y | Y | Y |
| Number of observations | 6,794 | 6,760 | 6,794 | 6,760 | 6,794 | 6,760 | 6,794 | 6,760 |
| Clusters | 108 | 108 | 108 | 108 | 108 | 108 | 108 | 108 |
| Adjusted $\mathrm{R}^{2}$ | 0.577 | 0.590 | 0.606 | 0.611 | 0.634 | 0.644 | 0.690 | 0.707 |

## Counterfactuals

1. Calculate the $h_{j t}$ chosen by a social planner who wants to maximize average learning:

- Average horizontal quality is .4.
- Small test score increases and larger ( $>20 \%$ ) decline in inequality.

2. Calculate the effects of choosing learning-maximizing $h_{j t}$ with improved sorting.

- Replace coefficients on value-added for rich and poor with $m \times \delta_{\text {rich }}$ for varying values of $m$.
- Get improvements relative to case where sorting improves but school quality is unchanged to isolate gains from changing match-specific quality.


## Counterfactual \#2

Changes in average value of $h_{j t}$


Why is this U-Shaped?

## Counterfactual \#2

Improvements in learning


## Counterfactual \#2

Improved sorting allows for product differentiation:


## Conclusion

- Instructional match has a large effect on learning.
- Substantial misallocation of instructional levels in rural Pakistan, a context much like other low-income countries.
- Competitive incentives mean that virtually all private schools compete to be the best school for richer students, providing the "wrong" type of instruction.
- Improving instructional match yields the largest benefits when students are more responsive to quality when they sort into schools.
- What's the policy implication?


## Drivers of Exit and Entry

|  | $(1)$ | $(2)$ | $(3)$ | $(4)$ | $(5)$ | $(6)$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1981 Population | 1998 Population | \% Pop. Change | Mean Assets | Percent Own Land | Gini Coefficient |
| num_privt | $248.879^{* *}$ | $373.904^{*}$ | 0.009 | 0.017 | -0.001 | 0.001 |
|  | $(108.864)$ | $(210.127)$ | $(0.025)$ | $(0.036)$ | $(0.010)$ | $(0.008)$ |
| Year FE | Y | Y | Y | Y | Y | Y |
| Village FE | N | N | N | Y | Y | Y |
| Number of observations | 448 | 448 | 436 | 448 | 336 | 448 |
| Clusters | 112 | 112 | 109 | 112 | 112 | 112 |
| Adjusted R $\mathrm{R}^{2}$ | 0.146 | 0.119 | -0.005 | 0.964 | 0.841 | 0.265 |

[^0]
## Relationship Between Type-Specific Value-Added Estimates



## Correlation Between Exit and Entry and Village Trends

|  | $(1)$ | $(2)$ | $(3)$ | $(4)$ | $(5)$ | $(6)$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1981 Population | 1998 Population | \% Pop. Change | Mean Assets | Percent Own Land | Gini Coefficient |
| num_privt | $248.879^{* *}$ | $373.904^{*}$ | 0.009 | 0.017 | -0.001 | 0.001 |
|  | $(108.864)$ | $(210.127)$ | $(0.025)$ | $(0.036)$ | $(0.010)$ | $(0.008)$ |
| Year FE | Y | Y | Y | Y | Y | Y |
| Village FE | N | N | N | Y | Y | Y |
| Number of observations | 448 | 448 | 436 | 448 | 336 | 448 |
| Clusters | 112 | 112 | 109 | 112 | 112 | 112 |
| Adjusted R $\mathrm{R}^{2}$ | 0.146 | 0.119 | -0.005 | 0.964 | 0.841 | 0.265 |

[^1]
## Demand Estimation: Step 2

Given estimates of $\zeta_{j n}$, use the relationship

$$
\zeta_{j n}=\xi_{j n}+\Gamma_{z}^{\text {school }} X_{j n}^{\text {school }}
$$

and GMM to estimate $\Gamma_{Z}^{s c h o o l}$ using residual variation in teacher salaries as an instrument for school fees. © Back

## Identifying Equilibrium Choice of Horizontal Quality

Divide $\frac{\partial s_{j t}}{\partial h_{j t}}$ by $\beta$ to get an identifying equation for each $j t$

$$
\begin{aligned}
& \sum_{i t} P\left(\text { type }_{i}=a d\right) \delta_{a d}\left(1-h_{j t}\right) p_{i j, a d, t}\left(1-p_{i j, a d, t}\right) \\
& \quad+\left(1-P\left(\text { type }_{i}=a d\right)\right) \delta_{d i s} h_{j t}\left(p_{i j, d i s, t}^{2}-p_{i j, d i s, t}\right)=0
\end{aligned}
$$

that does not contain any other unknowns but $h_{j t}$.

## Robustness: Peer and Class-size Controls

|  | $(1)$ <br> Math | $(2)$ <br> English | $(3)$ <br> Urdu | $(4)$ <br> Mean |
| :--- | :---: | :---: | :---: | :---: |
| $\mathbb{1}_{\text {ad }} \times$ num_prívt | $0.102^{*}$ | $0.078^{*}$ | 0.055 | $0.083^{* *}$ |
| num_privt | $(0.057)$ | $(0.044)$ | $(0.042)$ | $(0.036)$ |
|  | $-0.106^{* *}$ | -0.054 | -0.029 | $-0.064^{* *}$ |
| Peer Controls | $(0.036)$ | $(0.044)$ | $(0.035)$ | $(0.027)$ |
| Student-Teacher Ratio Controls | Y | Y | Y | Y |
| Lagged Test Score Controls | Y | Y | Y | Y |
| School by Type FE | Y | Y | Y | Y |
| Grade by Type FE | Y | Y | Y | Y |
| Year by Type FE | Y | Y | Y | Y |
| Number of observations | 6,788 | 6,788 | 6,788 | 6,788 |
| Clusters | 106 | 106 | 106 | 106 |
| Adjusted R ${ }^{2}$ | 0.583 | 0.602 | 0.640 | 0.703 |

[^2]
## Effects of Entry on Always-Private Students

|  | $(1)$ <br> Math | $(2)$ <br> English | $(3)$ <br> Urdu | $(4)$ <br> Mean |
| :--- | :---: | :---: | :---: | :---: |
| $\mathbb{1}_{\text {ad }} \times$ num_privt | $0.112^{*}$ | $0.101^{* *}$ | $0.098^{*}$ | $0.112^{* *}$ |
| num_pri ${ }_{\text {vt }}$ | $(0.067)$ | $(0.050)$ | $(0.050)$ | $(0.048)$ |
| Peer Controls | $-0.107^{* *}$ | -0.061 | -0.057 | $-0.078^{* *}$ |
| Number of Private Schools Controls | $(0.042)$ | $(0.058)$ | $(0.031)$ | $(0.034)$ |
| Lagged Test Score Controls | Y | Y | Y | Y |
| School by Type FE | Y | Y | Y | Y |
| Grade by Type FE | Y | Y | Y | Y |
| Year by Type FE | Y | Y | Y | Y |
| Number of observations | Y | Y | Y | Y |
| Clusters | 5,845 | 5,845 | 5,845 | 5,845 |
| Adjusted R2 | 108 | 108 | 108 | 108 |

## Descriptive Results on Determinants of School Choice

|  | $(1)$ <br> Changed <br> Schools | $(2)$ <br> Knows <br> Teacher's <br> Name | $(3)$ <br> Chose <br> School for <br> Distance | $(4)$ <br> Chose <br> School for <br> Quality | (5) <br> Mean <br> School <br> VA |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $P\left(\right.$ type $\left._{i}=a d\right)$ | $0.310^{* * *}$ | $0.261^{* * *}$ | -0.162 | $0.574^{* * *}$ | $0.153^{*}$ |
|  | $(0.059)$ | $(0.048)$ | $(0.104)$ | $(0.090)$ | $(0.081)$ |
| rank $_{i j}$ |  |  |  |  | $0.050^{* * *}$ |
|  |  |  |  |  | $(0.010)$ |
| $P\left(\right.$ type $\left._{i}=a d\right) \times$ rank $_{i j}$ |  |  |  |  | $0.043^{*}$ |
|  |  |  |  | $(0.023)$ |  |
| Mean | 0.347 | 0.532 | 0.427 | 0.210 | 0.022 |
| Observation Level | Child | Child-Year | Child | Child | Parent-School-Year |
| Number of observations | 5,621 | 13,645 | 2,873 | 2,873 | 22,826 |
| Clusters | 1,694 | 1,695 | 1,153 | 1,153 | 684 |
| Adjusted $\mathrm{R}^{2}$ | 0.008 | 0.005 | 0.002 | 0.038 | 0.031 |


[^0]:    Back

[^1]:    Back

[^2]:    Back

